

Causal effects in MPDAGs: identification and efficient estimation

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Goal

- Estimate the **total causal effect** of A on Y

Observational data

Randomized
control studies

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 - the change in Y due to $do(a)$ -
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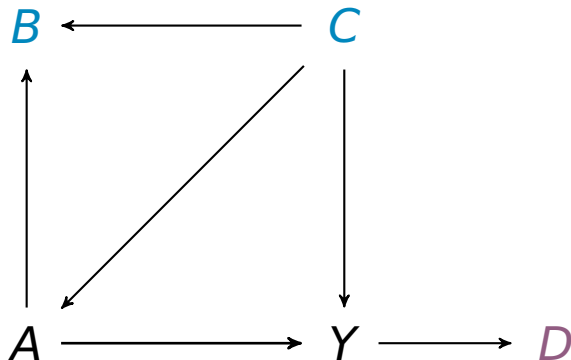
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 $f(y|do(a)) \neq f(y|a)$.

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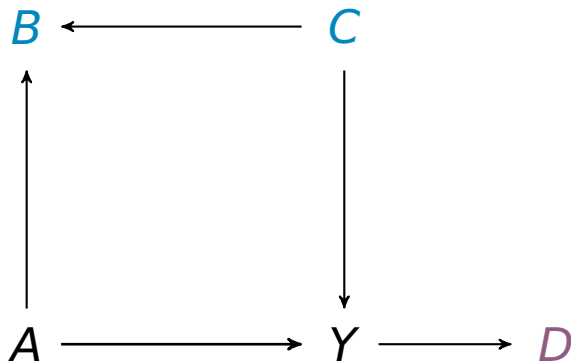
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Observational Causal DAG



Causal Directed Acyclic Graph (DAG) \mathcal{D} .

Interventional Causal DAG



Causal DAG \mathcal{D} after a “do”-intervention on A.

DAGs and distributions

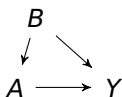
Interventional density

- $do(a)$: an intervention that sets variables A to a .
- Observational density $f(v)$, Interventional density $f(v|do(a))$.

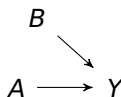
DAGs and distributions

Interventional density

- $do(a)$: an intervention that sets variables A to a .
- Observational density $f(v)$, Interventional density $f(v|do(a))$.
- A DAG \mathcal{D} is **causal** if for all observational and interventional densities:
$$f(v) = \prod_{V_j \in V} f(v_j | pa(v_j, \mathcal{D})) \quad \text{and} \quad f(v|do(a)) = \prod_{V_j \in V \setminus A} f(v_j | pa(v_j, \mathcal{D})).$$



$$f(b, a, y) = f(y|b, a)f(a|b)f(b)$$



$$f(b, y|do(a)) = f(y|b, a)f(b)$$

$$f(b, y|a) = f(y|b, a)f(b|a) \neq f(b, y|do(a))$$

How to define a causal effect?

Total causal effect

- Total causal effect - τ_{ay} - is some functional of $f(y|do(a))$, $P(Y|do(a))$.
- Examples: $E[Y|do(A = a + 1)] - E[Y|do(A = a)]$, $\frac{\partial}{\partial a}E(Y|do(a))$, OR, RR...

Identifiability

- A causal effect is identifiable from observational data if
 $f(y|do(a))$ is computable from $f(v)$.

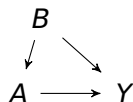
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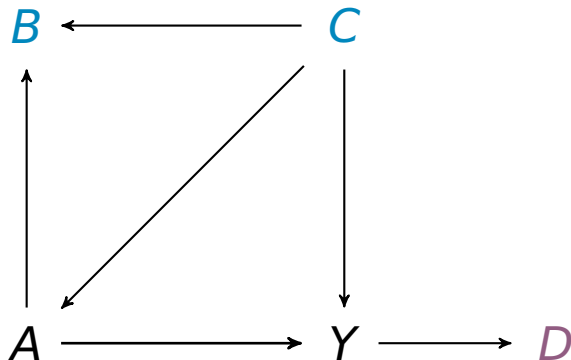
- A causal effect is **identifiable** from observational data if
 $f(y|do(a))$ is computable from $f(v)$.
- Given the causal DAG, every total causal effect is identifiable.



$$\begin{aligned}f(y|do(a)) &= \int f(b, y|do(a))db \\ &= \int f(y|b, a)f(b)db.\end{aligned}$$

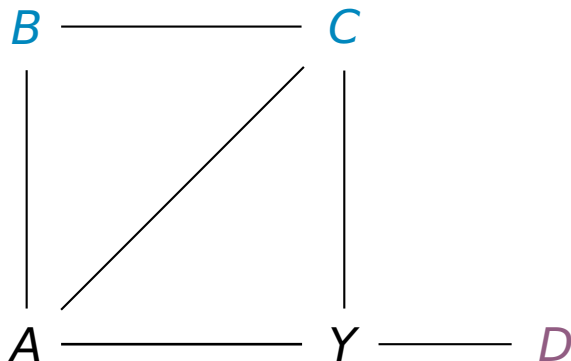
G-formula

Problem solved?



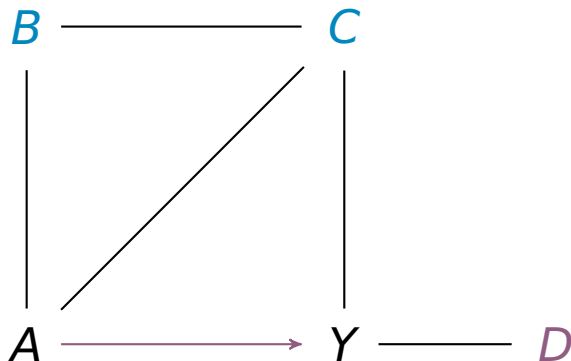
DAG \mathcal{D} .

Problem solved?



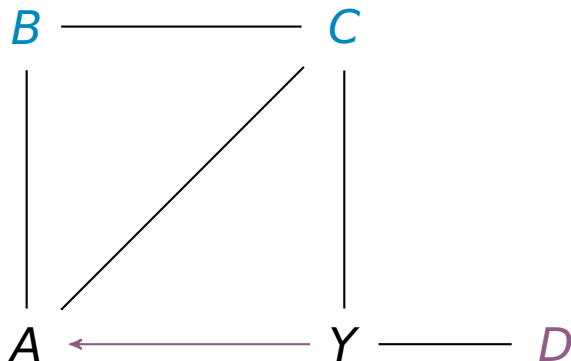
Completed Partially Directed Acyclic Graph (CPDAG) \mathcal{C} .

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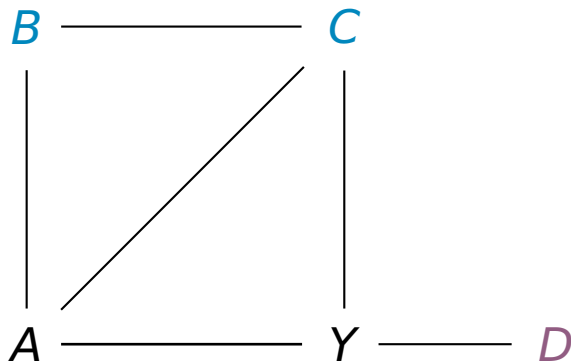
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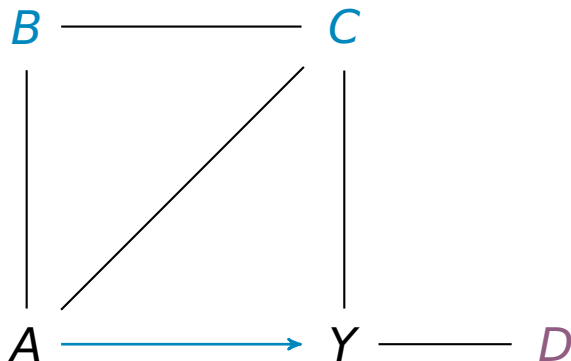
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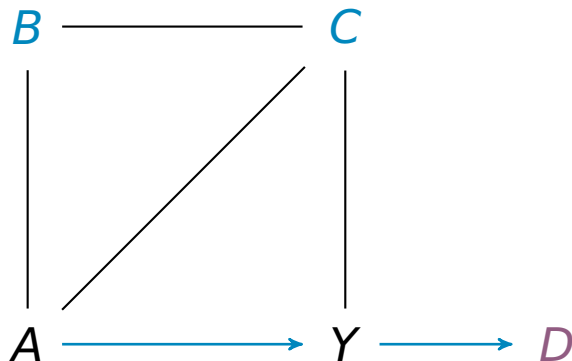
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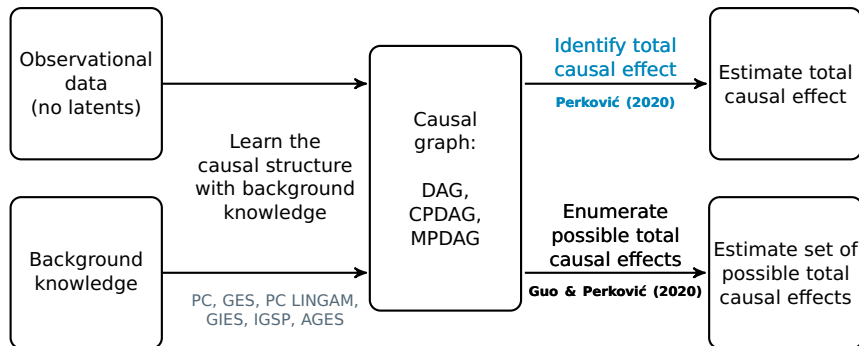
Partially Directed Acyclic Graph (PDAG).

Problem solved?



Maximally oriented Partially Directed Acyclic Graph (MPDAG) \mathcal{G} .

Framework

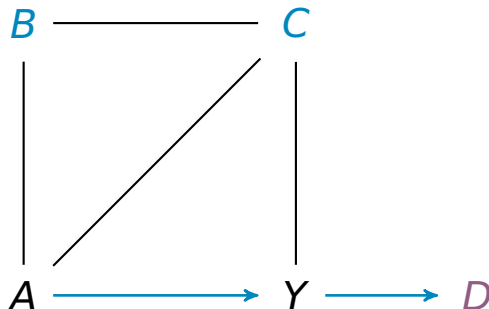


- PC (Spirtes et al, 1993), GES (Chickering, 2002)
- Adding background knowledge (Meek, 1995; TETRAD, Scheines et al., 1998), PC LINGAM (Hoyer et al., 2008), GIES (Hauser and Bühlmann, 2012), IGSP (Wang et al., 2017), etc.

Necessary and sufficient condition

Theorem (Perković, 2020)

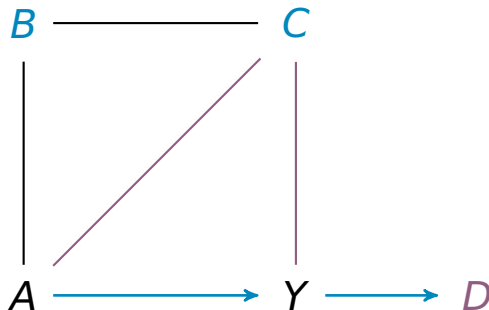
The total causal effect of A on Y is identifiable in MPDAG \mathcal{G} if and only if **all proper possibly causal paths** from A to Y start with a directed edge in \mathcal{G} .



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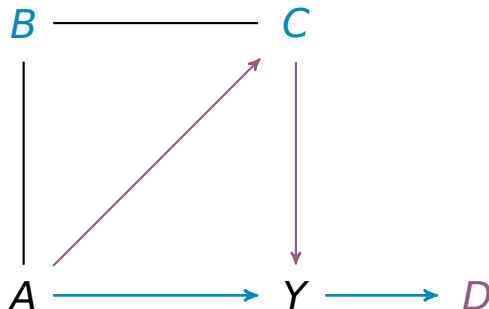
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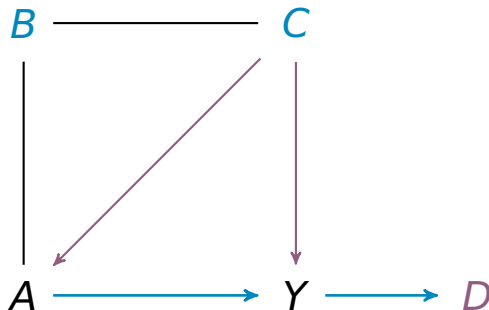
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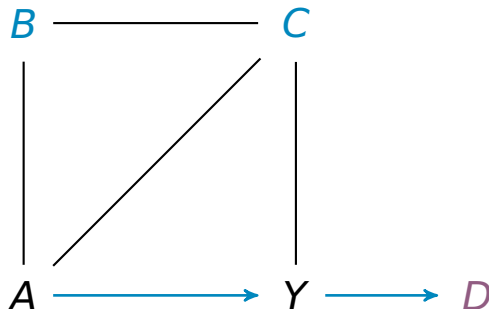
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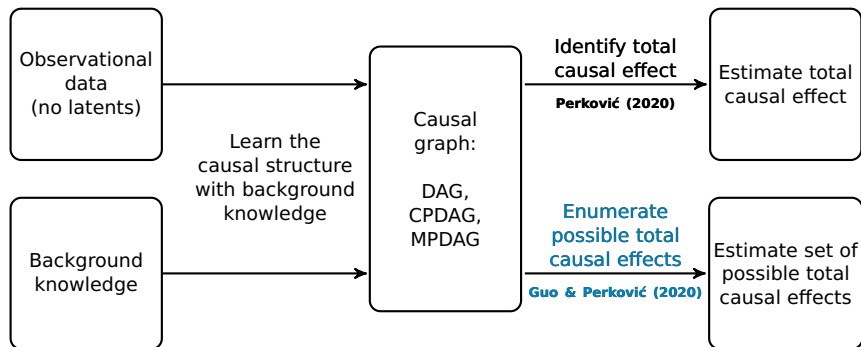
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Enumerate a set of possible total effects that contain the truth!

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- Naive: enumerate all the DAGs in $[\mathcal{G}]$, and identify for each DAG.
 - Computationally prohibitive for large $|V|$ (the complete graph contains $|V|!$ DAGs); see also Gillispie et al (2002), Steinsky et al (2013).

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- Enumerate the valid parent sets of A :
 - $|A|= 1$: IDA algorithm (Maathuis et al 2009).
 - $|A|> 1$: joint-IDA algorithm (Nandy et al 2017).

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 - Yet, the total effect and $f(y|do(a))$ can be the same for two different parent sets!

Optimal enumeration

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1. **complete:** $[\mathcal{G}] = [\mathcal{G}_1] \dot{\cup} [\mathcal{G}_2] \dot{\cup} \dots \dot{\cup} [\mathcal{G}_m]$
2. $f(y|\text{do}(a))$ is identifiable under each \mathcal{G}_i
3. **minimal:** maps $f \mapsto f(y|\text{do}(a))$ are distinct under each \mathcal{G}_i (identification formulae are distinct)
 \Rightarrow possible causal effects $f \mapsto \frac{\partial}{\partial a} \mathbb{E}(Y|\text{do}(A) = a)$ are distinct functionals!

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A natural algorithm is to recursively orient the undirected edges attached to A on **proper possibly causal paths** to Y .

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Input: MPDAG \mathcal{G} , $Y \in V$ and $A \subset V \setminus \{Y\}$.

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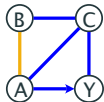
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MPDAG(\mathcal{G}, R) adds orientations R to \mathcal{G} and completes Meek orientation rules.

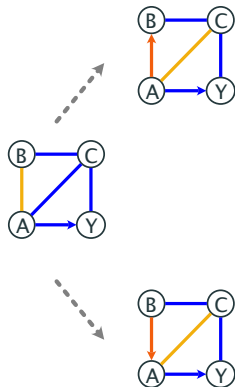
Optimal enumeration

Orienting $A - B$ first ...



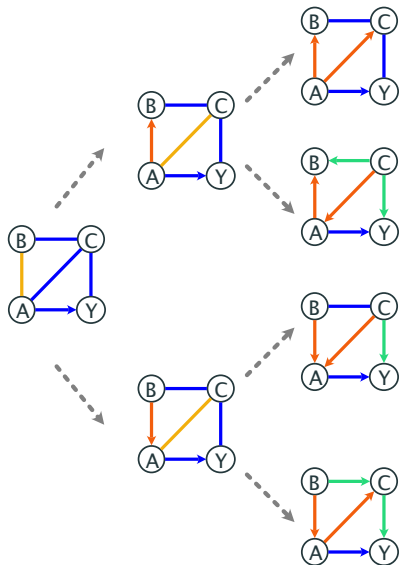
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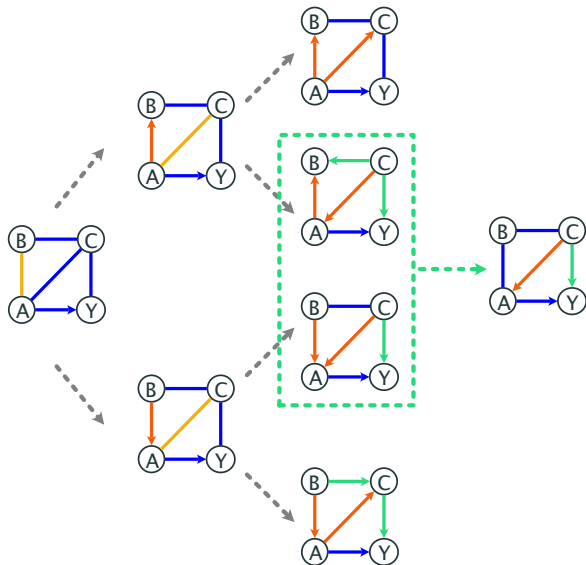
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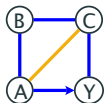
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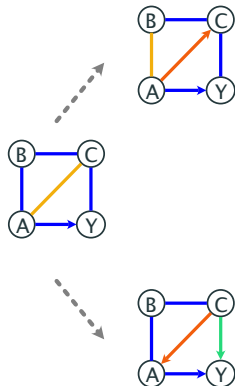
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Orienting $A - C$ first ...



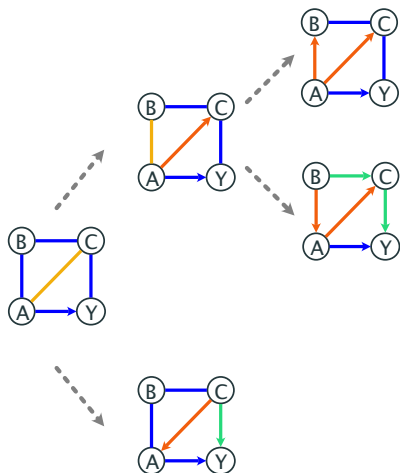
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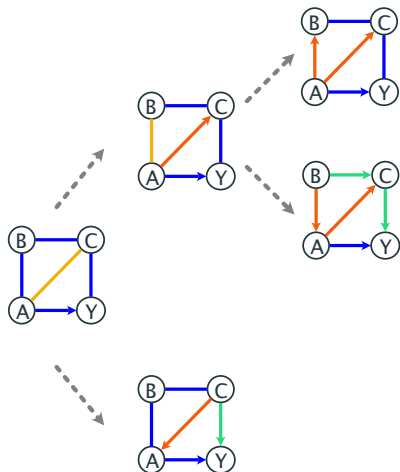
Orienting A – C first ...



Optimal enumeration

Orienting $A - C$ first ...

- $A - C$ should be oriented first because the *status* of $A - B - C - Y$ depends on $A - C - Y$.



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Algorithm IDGraphs, (Guo & Perković, 2020)

1. Pick $A_1 - V_1$ such that $A_1 \in A$ and A_1, V_1, \dots, Y is a **shortest** proper possibly causal path from A to Y .
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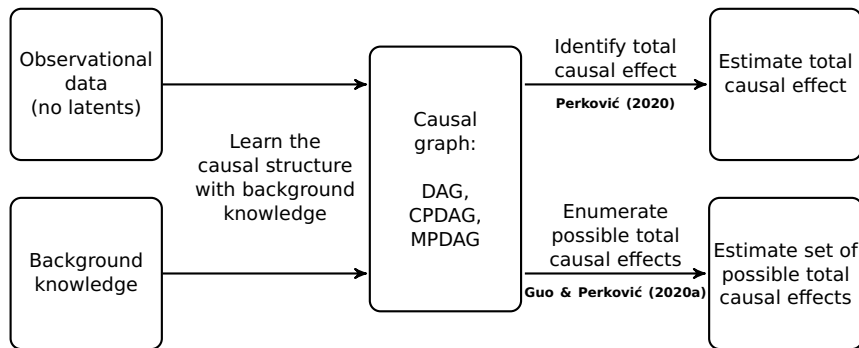
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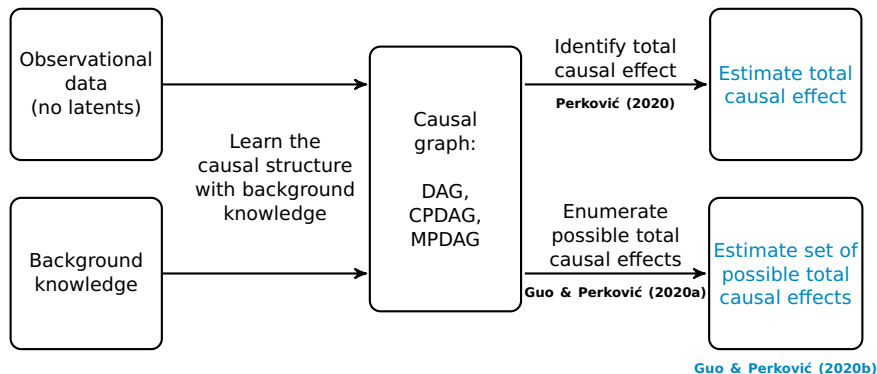
Hence, each \mathcal{G}_i represents the minimal set of additional orientations required for a particular interventional distribution/possible effect!

In contrast, the IDA algorithm will output 4 effects for this example, but two of them are different estimates of the same possible effect!

Framework

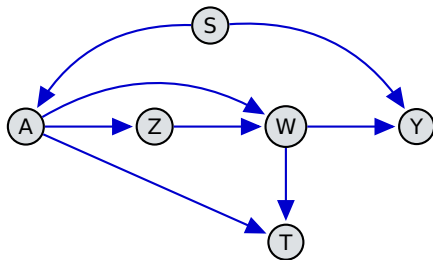


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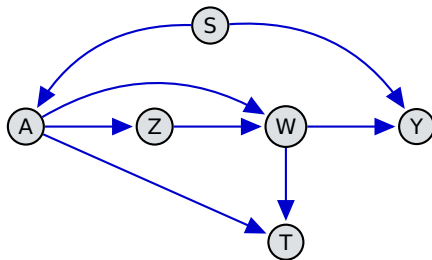


- In the following, we further assume **linearity** in the data generating mechanism.

Causal DAG, linear SEM

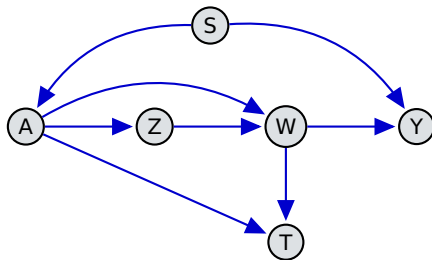


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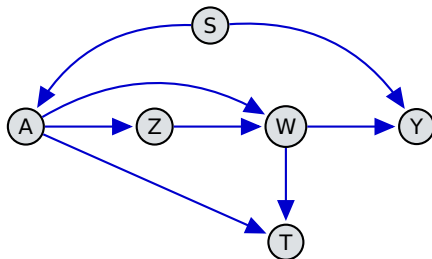
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Causal DAG, linear SEM



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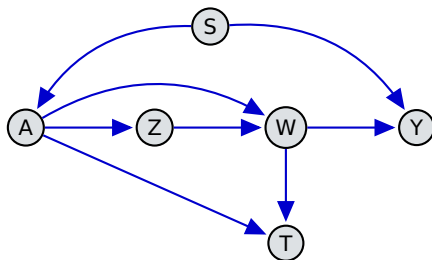
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- Suppose data is generated by a **linear structural equation model** (SEM)

$$X_V = \sum_{u:u \rightarrow V} \gamma_{uv} X_u + \epsilon_u, \quad \mathbb{E} \epsilon_u = 0, \quad 0 < \text{var} \epsilon_u < \infty.$$

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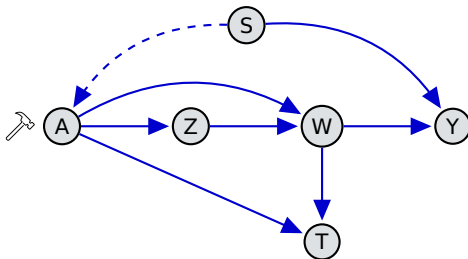
- Under no unobserved confounder, the errors are **mutually independent**.

Total effect

Suppose we want to estimate the **total (causal) effect of A on Y**.

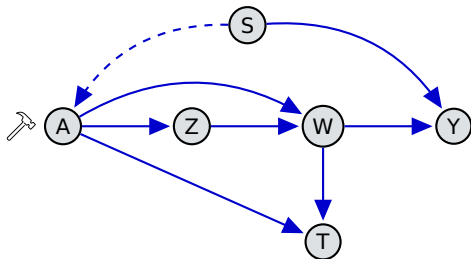
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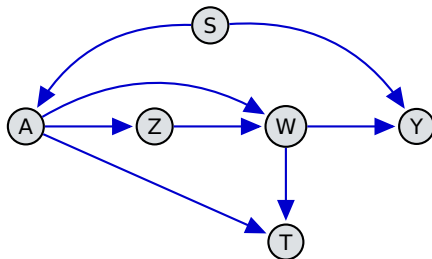
The total effect τ_{AY} is defined as the slope of $x_a \mapsto \mathbb{E}[X_Y | \text{do}(X_A = x_a)]$, given by a sum-product of Wright (1934):

$$\tau_{AY} = \frac{d}{dx_a} \mathbb{E}[X_Y | \text{do}(X_A = x_a)] = (\gamma_{AZ}\gamma_{ZW} + \gamma_{AW})\gamma_{WY}.$$

Estimation

Our task is to estimate τ_{AY} from n iid observational sample generated by a linear SEM associated with causal DAG \mathcal{D} , given that

$\mathcal{D} \in [\mathcal{G}]$ for MPDAG \mathcal{G} , τ_{AY} is identifiable from \mathcal{G} .

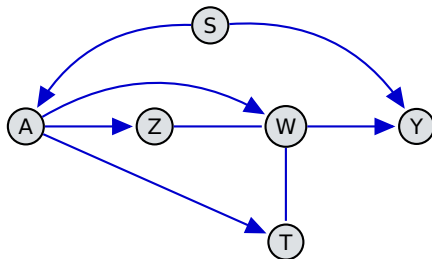


DAG \mathcal{D}

Estimation

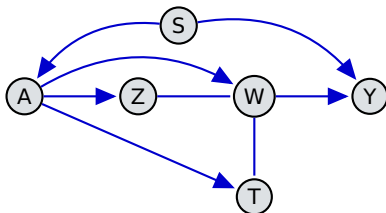
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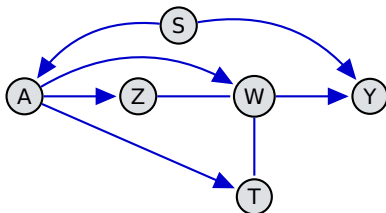


MPDAG \mathcal{G}

Buckets

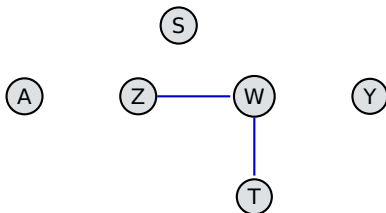


Buckets



- Let “buckets” be the undirected connected components of MPDAG \mathcal{G} :

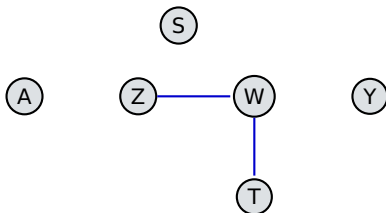
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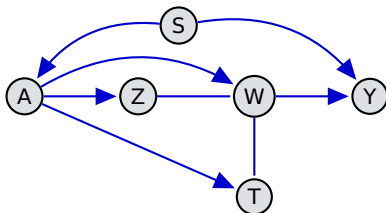


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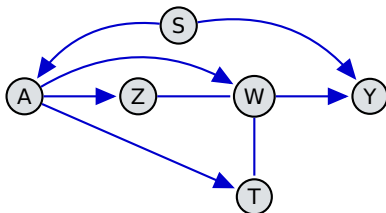


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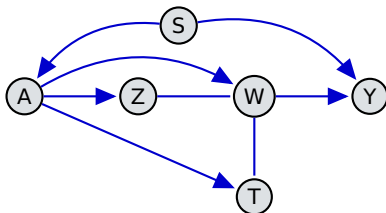


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- We use this to reparametrize the SEM.

Block-recursive reparametrization

Proposition (Block-recursive form, Guo and Perković, 2020)

Let B_1, \dots, B_K be the ordered bucket decomposition of V in MPDAG \mathcal{G} . Then

$$\begin{aligned} X &= \Lambda^T X + \varepsilon, & \Lambda &= (\lambda_{ij}), j \in B_k, i \notin \text{pa}(B_k, \mathcal{G}) \Rightarrow \lambda_{ij} = 0, \\ \mathbb{E} \varepsilon &= \mathbf{0}, & \mathbb{E} \varepsilon_{B_k} \varepsilon_{B_k}^T &\succ \mathbf{0}, \quad \varepsilon_{B_k} \text{ mutually independent,} \end{aligned}$$

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$$\tau_{AY} = \Lambda_{A,S} \left[(I - \Lambda_{S,S})^{-1} \right]_{S,Y}.$$

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 - This is a special case of “seemingly unrelated regression” under the **restrictive property**.

Efficient estimator

If τ_{AY} is identifiable given MPDAG \mathcal{G} , the **\mathcal{G} -regression estimator** is defined as:

$$\hat{\tau}_{AY}^{\mathcal{G}} := \hat{\Lambda}_{A,S}^{\mathcal{G}} \left[(I - \hat{\Lambda}_{S,S}^{\mathcal{G}})^{-1} \right]_{S,Y},$$

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Theorem (\mathcal{G} -regression, Guo and Perković, 2020)

Then for any regular estimator $\hat{\tau}_{AY}$ that only uses the **first two moments** of the data, it holds that

$$\text{avar}(\hat{\tau}_{AY}) \geq \text{avar}\left(\hat{\tau}_{AY}^{\mathcal{G}}\right).$$

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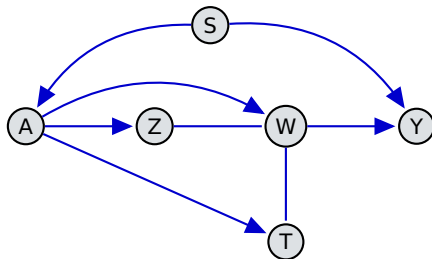
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This includes estimators in the literature:

- covariate adjustment (Henckel et al., 2019, Witte et al., 2020),
- recursive regressions (Nandy et al., 2017, Gupta et al., 2020),
- modified Cholesky decomposition (Nandy et al., 2017).

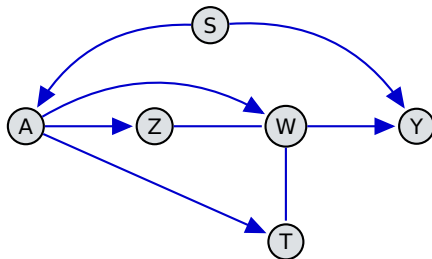
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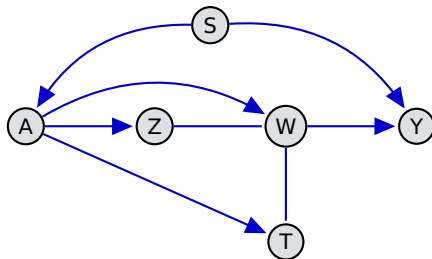
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where $\hat{\lambda}_{AW}$, $\hat{\lambda}_{WY}$ are taken from $W \sim A$ and $Y \sim W + S$ respectively.

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\mathcal{G} -regression estimator

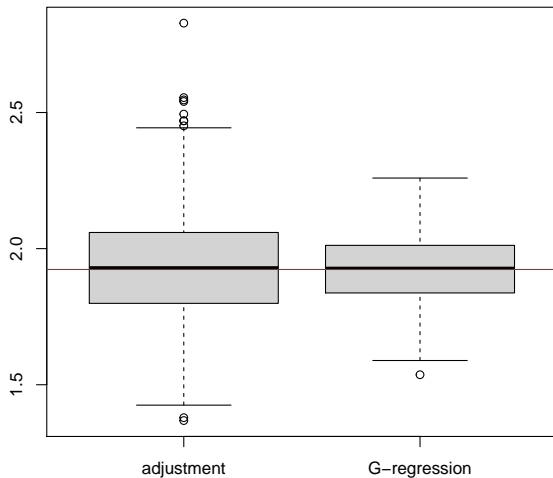
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adjustment estimator

$$\hat{\tau}_{AY}^{\text{adj}} = \hat{\beta}_{AY} \text{ from } Y \sim A + S.$$

Efficient estimator



$n = 100, t_5$ errors.

Simulation: relative efficiency

Table: Geometric average of squared errors relative to \mathcal{G} -regression:

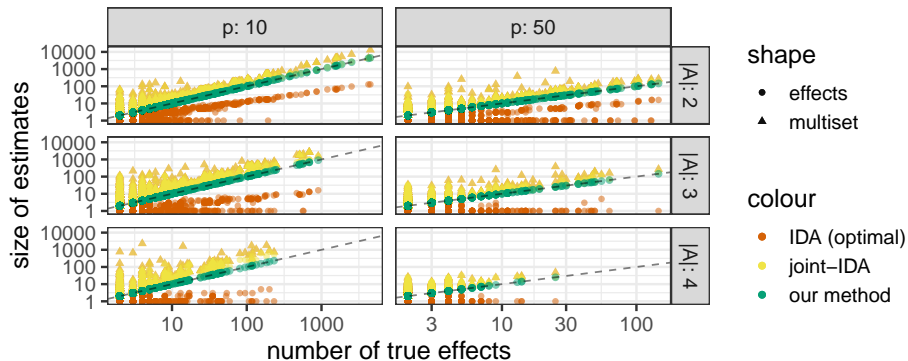
adj.0: optimal adjustment

IDA.M: IDA (Cholesky)

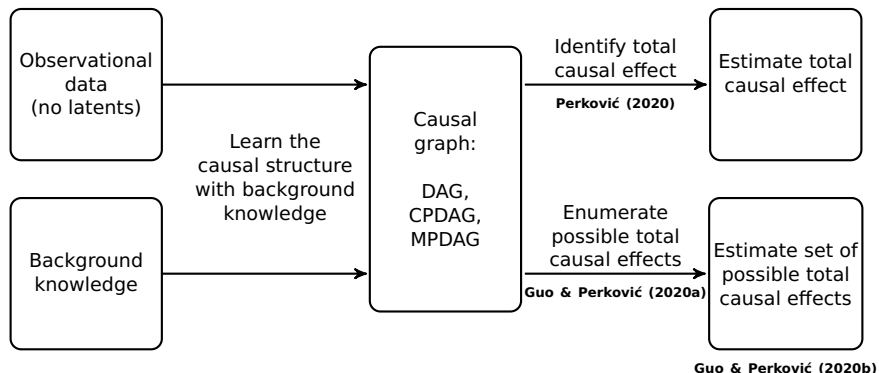
IDA.R: IDA (recursive regression)

| $ A $ | $ V =20$ | | $ V =50$ | | $ V =100$ | |
|-------|----------|----------|----------|----------|-----------|----------|
| | $n=100$ | $n=1000$ | $n=100$ | $n=1000$ | $n=100$ | $n=1000$ |
| adj.0 | | | | | | |
| 1 | 1.3 | 1.3 | 1.4 | 1.3 | 1.5 | 1.5 |
| 2 | 3.4 | 4.2 | 4.7 | 4.9 | 4.2 | 4.5 |
| 3 | 6.3 | 5.9 | 7.4 | 7.2 | 7.8 | 8.0 |
| 4 | 9.3 | 9.3 | 12 | 14 | 12 | 12 |
| IDA.M | | | | | | |
| 1 | 20 | 19 | 61 | 48 | 103 | 108 |
| 2 | 62 | 65 | 220 | 182 | 293 | 356 |
| 3 | 93 | 119 | 354 | 396 | 749 | 771 |
| 4 | 154 | 222 | 533 | 895 | 1188 | 1604 |
| IDA.R | | | | | | |
| 1 | 20 | 19 | 61 | 48 | 103 | 108 |
| 2 | 33 | 38 | 121 | 113 | 176 | 199 |
| 3 | 30 | 39 | 171 | 135 | 342 | 312 |
| 4 | 48 | 50 | 187 | 214 | 405 | 432 |

Simulation: size of possible effects



Final remarks



- **R package** `eff2`: github.com/richardkwo/eff2
- **Efficient estimation beyond linear SEMs?** We are working on it!

Thanks!